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Errata


From: RAND Corporation Publishing Services

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Re: Corrected pages (pp. 30, 100)

The total Air Force population on pages 30 and 100 was erroneously reported as 328,667. The correct number in both instances is 504,414. This error does not affect the authors' analysis or conclusions.

We regret any inconvenience.
The mean score for problematic Internet use among Airmen was 26 out of a possible 75. This average was similar to the score of 33 found for a nonrandom U.S. sample of 785 people, which is not necessarily a comparable population (Caplan, 2010). Figure 5.3 shows the Airman scores on the GPIUS2. Higher scores indicate greater agreement with the survey items shown in Table 5.5 and thus greater problems with Internet use. The lowest scores indicate little or no Internet use that is poorly-controlled and problematic. While most Airmen fall in the lower ranges of GPIUS2 scores, about 6 percent fall into the top (worst) ranges. If this survey population can be generalized to the total Air Force military population (504,414 as of February 2012), then more than 30,000 Airmen may struggle with problematic Internet use.

**Figure 5.3. Airman Problematic Internet Use Scores**

![Bar chart showing Airman Problematic Internet Use Scores](image)

NOTE: N = 3,393.

What this scale does not capture is Airman activities online. Possible compulsive activities include gaming, gambling, shopping, web surfing, online sexual activity, and chatting with others. Additional research would be necessary to identify whether Airmen with problematic Internet use tend to be engaged in any particular type of activity. The scale also does not capture problematic use of other ICTs, such as texting and mobile phone applications.

**Airman Problematic Internet Use Is Linked to Well-Being**

Airman scores on the GPIUS2 were significantly correlated ($p < 0.0001$) with several measures of their well-being. In particular, as indications of problematic Internet use increased, self-rated mental health decreased ($r = 0.25$), depressed mood increased ($r = 0.18$), and loneliness increased ($r = 0.33$). These associations do not necessarily suggest that problematic

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2 Recent prevalence estimates for problematic Internet use among the adult U.S. population are not available.
similar across neighboring cells. For this reason, we decided to collapse the cell of active-duty female officers aged 18 to 24 with the cell of active-duty female enlisted Airmen aged 18 to 24. More generally, we collapsed the two rank cells defined by crossing affiliation and gender within the 18–24 age category. The rest of the cells were untouched. Collapsing modifies the post-stratification weights for only those cells that are collapsed.

If \( i \) and \( j \) are the two cells to be collapsed, the post-stratification weight for those subjects that belong to either cell \( i \) or \( j \) is given by \( w_{i,j} = (N_i+N_j) / (r_i+r_j) \). The weights for the other cells \( h \) are equal to \( w_h \) as defined above.

Collapsing should improve the accuracy of the survey estimates and obtain a less-biased age distribution in the sample.

We weighted our sample of active-duty Airmen to ensure that the proportions in it were demographically representative of the larger Air Force population \((N = 504,414)\). We used the following demographic categories:

- Age group (18–24, 25–34, 35–44, 45 and older)
- Gender (men, women)
- Rank group (enlisted, officer).

The expected distribution of the respondents across these three demographic categories would be that of a random sample of the active-duty component. We therefore weighted the proportions of respondents to those of the entire Air Force active-duty population, using this expected distribution (see Table C.2).

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2 We used Air Force population data for 17- to 24-year-old Airmen to estimate the proportion of 18- to 24-year-old Airmen in our sample. For human subjects protection purposes, we excluded 17-year-olds before drawing the random sample of invitees.